

## Randomised Algorithms

Lecture 2: Concentration Inequalities, Application to Balls-into-Bins

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## Outline

How to Derive Chernoff Bounds

Application 1: Balls into Bins

Appendix: More on Moment Generating Functions (non-examinable)

## General Recipe for Deriving Chernoff Bounds

Recipe

The **three main steps** in deriving Chernoff bounds for sums of **independent** random variables  $X = X_1 + \dots + X_n$  are:

1. Instead of working with  $X$ , we switch to the **moment generating function**  $e^{\lambda X}$ ,  $\lambda > 0$  and apply Markov's inequality  $\sim \mathbf{E}[e^{\lambda X}]$
2. Compute an upper bound for  $\mathbf{E}[e^{\lambda X}]$  (using independence)
3. Optimise value of  $\lambda$  to obtain best tail bound

## Chernoff Bound: Proof

Chernoff Bound (General Form, Upper Tail)

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then, for any  $\delta > 0$  it holds that

$$\mathbf{P}[X \geq (1 + \delta)\mu] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1 + \delta)}} \right]^\mu.$$

Proof:

1. For  $\lambda > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)\mu] \stackrel{e^{\lambda X} \text{ is incr}}{=} \mathbf{P}[e^{\lambda X} \geq e^{\lambda(1 + \delta)\mu}] \stackrel{\text{Markov}}{\leq} e^{-\lambda(1 + \delta)\mu} \mathbf{E}[e^{\lambda X}]$$

$$2. \mathbf{E}[e^{\lambda X}] = \mathbf{E}[e^{\lambda \sum_{i=1}^n X_i}] \stackrel{\text{indep}}{=} \prod_{i=1}^n \mathbf{E}[e^{\lambda X_i}]$$

- 3.

$$\mathbf{E}[e^{\lambda X_i}] = e^{\lambda p_i} + (1 - p_i) = 1 + p_i(e^\lambda - 1) \stackrel{1 + x \leq e^x}{\leq} e^{p_i(e^\lambda - 1)}$$

## Chernoff Bound: Proof

1. For  $\lambda > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)\mu] \stackrel{\substack{e^{\lambda x} \text{ is incr} \\ \text{Markov}}}{=} \mathbf{P}[e^{\lambda X} \geq e^{\lambda(1+\delta)\mu}] \leq e^{-\lambda(1+\delta)\mu} \mathbf{E}[e^{\lambda X}]$$

2.  $\mathbf{E}[e^{\lambda X}] = \mathbf{E}[e^{\lambda \sum_{i=1}^n X_i}] \stackrel{\text{indep}}{=} \prod_{i=1}^n \mathbf{E}[e^{\lambda X_i}]$

3.  $\mathbf{E}[e^{\lambda X_i}] = e^{\lambda p_i} + (1 - p_i) = 1 + p_i(e^\lambda - 1) \stackrel{1+x \leq e^x}{\leq} e^{p_i(e^\lambda - 1)}$

4. Putting all together

$$\mathbf{P}[X \geq (1 + \delta)\mu] \leq e^{-\lambda(1+\delta)\mu} \prod_{i=1}^n e^{p_i(e^\lambda - 1)} = e^{-\lambda(1+\delta)\mu} e^{\mu(e^\lambda - 1)}$$

5. Choose  $\lambda = \log(1 + \delta) > 0$  to get the result.

## Chernoff Bounds: Lower Tails

We can also use Chernoff Bounds to show a random variable is **not too small** compared to its mean:

Chernoff Bounds (General Form, Lower Tail)

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then, for any  $0 < \delta < 1$  it holds that

$$\mathbf{P}[X \leq (1 - \delta)\mu] \leq \left[ \frac{e^{-\delta}}{(1 - \delta)^{1-\delta}} \right]^\mu,$$

and thus, by substitution, for any  $t < \mu$ ,

$$\mathbf{P}[X \leq t] \leq e^{-\mu} \left( \frac{e\mu}{t} \right)^t.$$

**Exercise on Supervision Sheet**

Hint: multiply both sides by  $-1$  and repeat the proof of the Chernoff Bound

## Nicer Chernoff Bounds

“Nicer” Chernoff Bounds

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then,

▪ For all  $t > 0$ ,

$$\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$$

$$\mathbf{P}[X \leq \mathbf{E}[X] - t] \leq e^{-2t^2/n}$$

▪ For  $0 < \delta < 1$ ,

$$\mathbf{P}[X \geq (1 + \delta)\mathbf{E}[X]] \leq \exp\left(-\frac{\delta^2 \mathbf{E}[X]}{3}\right)$$

$$\mathbf{P}[X \leq (1 - \delta)\mathbf{E}[X]] \leq \exp\left(-\frac{\delta^2 \mathbf{E}[X]}{2}\right)$$

All upper tail bounds hold even under a **relaxed independence assumption**:  
For all  $1 \leq i \leq n$  and  $x_1, x_2, \dots, x_{i-1} \in \{0, 1\}$ ,

$$\mathbf{P}[X_i = 1 \mid X_1 = x_1, \dots, X_{i-1} = x_{i-1}] \leq p_i.$$

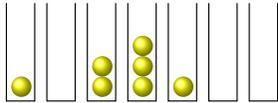
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## Balls into Bins



Balls into Bins Model

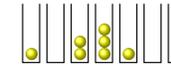
You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked **independently and uniformly at random**.

- A very natural but also rich **mathematical** model
- In **computer science**, there are several interpretations:
  1. Bins are a hash table, balls are items
  2. Bins are processors and balls are jobs
  3. Bins are data servers and balls are queries



**Exercise:** Think about the relation between the **Balls into Bins Model** and the **Coupon Collector Problem**.

## Balls into Bins: Bounding the Maximum Load (1/4)



Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked **independently and uniformly at random**.

**Question 1:** How large is the **maximum load** if  $m = 2n \log n$ ?

- Focus on an **arbitrary single** bin. Let  $X_i$  the indicator variable which is 1 iff ball  $i$  is assigned to this bin. Note that  $p_i = \mathbf{P}[X_i = 1] = 1/n$ .
- The total balls in the bin is given by  $X := \sum_{i=1}^m X_i$ . here we could have used the “nicer” bounds as well!
- Since  $m = 2n \log n$ , then  $\mu = \mathbf{E}[X] = 2 \log n$

$$\mathbf{P}[X \geq t] \leq e^{-\mu} (e\mu/t)^t$$

- By the Chernoff Bound,

$$\mathbf{P}[X \geq 6 \log n] \leq e^{-2 \log n} \left( \frac{2e \log n}{6 \log n} \right)^{6 \log n} \leq e^{-2 \log n} = n^{-2}$$

## Balls into Bins: Bounding the Maximum Load (2/4)

- Let  $\mathcal{E}_j := \{X(j) \geq 6 \log n\}$ , that is, bin  $j$  receives at least  $6 \log n$  balls.
- We are interested in the probability that **at least** one bin receives at least  $6 \log n$  balls  $\Rightarrow$  this is the event  $\bigcup_{j=1}^n \mathcal{E}_j$
- By the **Union Bound**,

$$\mathbf{P}\left[\bigcup_{j=1}^n \mathcal{E}_j\right] \leq \sum_{j=1}^n \mathbf{P}[\mathcal{E}_j] \leq n \cdot n^{-2} = n^{-1}.$$

- Therefore **whp**, no bin receives at least  $6 \log n$  balls
- By **pigeonhole principle**, the max loaded bin receives at least  $2 \log n$  balls. Hence our bound is pretty sharp.

**whp** stands for *with high probability*:

An event  $\mathcal{E}$  (that implicitly depends on an input parameter  $n$ ) occurs **whp** if

$$\mathbf{P}[\mathcal{E}] \rightarrow 1 \text{ as } n \rightarrow \infty.$$

This is a very standard notation in randomised algorithms but it may vary from author to author. **Be careful!**

## Balls into Bins: Bounding the Maximum Load (3/4)

**Question 2:** How large is the **maximum load** if  $m = n$ ?

- Using the **Chernoff Bound**:  $\mathbf{P}[X \geq t] \leq e^{-\mu} (e\mu/t)^t$

$$\mathbf{P}[X \geq t] \leq e^{-1} \left(\frac{e}{t}\right)^t \leq \left(\frac{e}{t}\right)^t$$

- By setting  $t = 4 \log n / \log \log n$ , we claim to obtain  $\mathbf{P}[X \geq t] \leq n^{-2}$ .
- Indeed:

$$\left(\frac{e \log \log n}{4 \log n}\right)^{4 \log n / \log \log n} = \exp\left(\frac{4 \log n}{\log \log n} \cdot \log\left(\frac{e \log \log n}{4 \log n}\right)\right)$$

- The term inside the exponential is

$$\frac{4 \log n}{\log \log n} \cdot (\log(e/4) + \log \log \log n - \log \log n) \leq \frac{4 \log n}{\log \log n} \left(-\frac{1}{2} \log \log n\right),$$

obtaining that  $\mathbf{P}[X \geq t] \leq n^{-4/2} = n^{-2}$ .

This inequality only works for large enough  $n$ .

## Balls into Bins: Bounding the Maximum Load (4/4)

We just proved that

$$\mathbf{P}[X \geq 4 \log n / \log \log n] \leq n^{-2},$$

thus by the **Union Bound**, no bin receives more than  $\Omega(\log n / \log \log n)$  balls with probability at least  $1 - 1/n$ .  $\square$

- One can prove that **whp** at least one bin receives at least  $c \log n / \log \log n$  balls, for some constant  $c > 0$ .

## Conclusions

- If the number of balls is  $2 \log n$  times  $n$  (the number of bins), then to distribute balls at random is a **good algorithm**
  - This is because the worst case maximum load is whp.  $6 \log n$ , while the average load is  $2 \log n$
- For the case  $m = n$ , the algorithm is **not good**, since the maximum load is whp.  $\Theta(\log n / \log \log n)$ , while the average load is 1.

### A Better Load Balancing Approach

For any  $m \geq n$ , we can improve this by sampling **two bins** in each step and then assign the ball into the bin with lesser load.  
 $\Rightarrow$  for  $m = n$  this gives a maximum load of  $\log_2 \log n + \Theta(1)$  w.p.  $1 - 1/n$ .

This is called the **power of two choices**: It is a common technique to improve the performance of randomised algorithms (covered in Chapter 17 of the textbook by Mitzenmacher and Upfal)

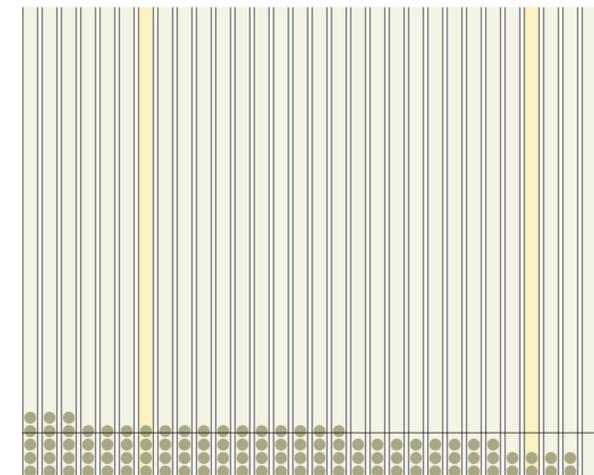
## ACM Paris Kanellakis Theory and Practice Award 2020



For “the discovery and analysis of balanced allocations, known as the *power of two choices*, and their extensive applications to practice.”

“These include *i-Google’s web index*, *Akamai’s overlay routing network*, and highly reliable *distributed data storage systems* used by *Microsoft and Dropbox*, which are all based on variants of the *power of two choices paradigm*. There are many other software systems that use balanced allocations as an important ingredient.”

## Simulation



Sampled two bins u.a.r.

Next Step Advance by 60 Go Trim Interval (ms): 1 Sort in each round Auto-trim Draw mean  
Number of bins: 3 Capacity: 3 Reset Process: Two-Choice Batch size: 9 Noise (g): 5  
Plot: (MAX-NORMALISED-LOAD) Add Initialise configuration: (EMPTY) Init

[https://www.dimitrioslos.com/balls\\_and\\_bins/visualiser.html](https://www.dimitrioslos.com/balls_and_bins/visualiser.html)

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## Moment Generating Functions (non-examinable)

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Moment-Generating Function

The **moment-generating** function of a random variable  $X$  is

$$M_X(t) = \mathbf{E} \left[ e^{tX} \right], \quad \text{where } t \in \mathbb{R}.$$

Using power series of  $e$  and differentiating shows that  $M_X(t)$  encapsulates all moments of  $X$ .

Lemma

1. If  $X$  and  $Y$  are two r.v.'s with  $M_X(t) = M_Y(t)$  for all  $t \in (-\delta, +\delta)$  for some  $\delta > 0$ , then the distributions  $X$  and  $Y$  are identical.
2. If  $X$  and  $Y$  are **independent** random variables, then

$$M_{X+Y}(t) = M_X(t) \cdot M_Y(t).$$

Proof of 2:

$$M_{X+Y}(t) = \mathbf{E} \left[ e^{t(X+Y)} \right] = \mathbf{E} \left[ e^{tX} \cdot e^{tY} \right] \stackrel{(!)}{=} \mathbf{E} \left[ e^{tX} \right] \cdot \mathbf{E} \left[ e^{tY} \right] = M_X(t) M_Y(t) \quad \square$$